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VI.	CALL FOR PAPERS

A Comparative Analysis of Books

on Multiple Regression

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带了"你们,你就是我愿意的吗?" 网络拉马拉马拉马拉

Dennis W. Leitner

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Randall E. Schumacker

Southern Illinois University at Carbondale

ABSTRACT

The purpose of this article was to <u>compare 33 books</u> on multiple regression that have appeared in past years. The books were compared on topics covered; year and publisher; orientation; level of presentation; background required; published reviews; and readability. This information should aid students and instructors in selecting a multiple regression book for a given topic, approach, or need.

INTRODUCTION

In the past decade and a half, many books and reference works on the topic of multiple regression have become available. This is probably due to a combination of several factors, some of which might be the following: (1) The widespread availability of computers and computer programs. (2) The pioneering work of Robert Bottenberg, Joe Mard, Earl Jennings and their disciples from Texas, and (3) The excellent articles by Jacob Cohen entitled "Multiple regression as a general data-analytic device" and Richard Darlington entitled "Multiple regression in psychological and research practice", both of which appeared in <u>Psychological Bulletin</u>. The books were compared on the following information: (1) topics covered; (2) year and publisher; (3) orientation (theoretical vs applied); (4) <u>level</u> of presentation (textbook vs reference); (5) background required (basic statistics vs matrix algebra); (6) reference list of published reviews; and (7) readability.

MULTIPLE LINEAR REGRESSION VIEWPOINTS

"Chairman	Joe Ward 167 East Arrowhead Drive San Antonio, TX 78228
Editor	The University of Akron Akron, OH 44325
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Ellinger (1981) presented a historical overview of readability formulae

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from 1921 to the present. She included the criteria on which they were based, factors evaluated, reliability and validity. The authors selected

the Flesch Ease Formula (Flesch, 1948) because it is appropriate for adult reading material and does not require use of a word list. The Flesch Ease

Formula is:

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Score = 206.835 - (...846 + NSYL) - (...1.015 + SL)

NSYL = number of syllables per a 100 word sample where: "这个人,不是这些好,我们没有这些的人,就是这些你的人,就是一些你的你的你都是你。" SL = average sentence length

12. **XE** 25 : The reading ease score can range from 0 to 100 with a lower score indicating afield low man advected day provering and region with the entropy of a providence a more difficult reading level (see Table 3). My set of the set of the set 化物品 医皮上皮 无限的 化化化物化合物 化化化合物 化化合物化合物合物 医鼻前外 化合物合物 化化合物 化分析 化分析 1944年1月19日19月1日日本,世界推荐:「安安日本,直有27月2日年期,隆江道,了马**城,和**国国际部队,不是公司了一个一个工作的 and Antherapping and the production of the

Book Selection

The books selected for comparison represent those from the social sciences and statistical literature that an educational researcher might be inclined to use. The list is by no means exhaustive. The books are listed alphabetically by author in a separate section of the references. and the contract of and have a cash in the

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Criteria for Topic Inclusion

资源公式资产集中,中部公共。 The books were compared on selected topics that the authors felt readers would be interested in. A particular topic was indicated as being covered in a given book if sufficient presentation or explanation was present, e.g. an entry in the table of contents, subject index, or several pages of discussion. For example, linear regression meant the straight-line one predictor variable regression model.

Anova compared to regression implied a comparison of the two as special cases of the general linear model. Matrix algebra referred to whether or not the book contained matrix algebra computations or had an appendix with matrix algebra rules and procedures. The covariance topic included the discussion of analysis of covariance. The other selected topics are selfexplanatory.

RESULTS

Table 1 indicates a comparison of topics among the books. Most books covered the following topics: linear regression, curvilinear, polynomial or nonlinear regression, zero-order correlation, matrix algebra, partial/semipartial correlation, and anova compared to regression, respectively.

Several books published between 1966 and 1976 had included <u>computer</u> programs while later books did not because of the popularity of several statistical packages. For example, Younger (1979) published a book which included comparisons among SAS, SPSS, and BMDP statistical packages in performing numerous applied regression examples.

Table 2 indicates a comparison among the books according to year, publisher, orientation, presentation, and background. Most books appeared in the 1970's (1960's, n=9; 1970's, n=18; 1980's, n=6). The majority had an applied orientation (applied, n=21; theoretical, n=9; both, n=3) with both a textbook and reference level of presentation (textbook, n=9; reference, n=12; both, n=12). Most books also required a basic statistics background (basic statistics, n=21; matrix algebra, n=12). Books with a theoretical orientation usually required knowledge of matrix algebra.

<u>Published reviews were</u> found for many of the books. These are listed alphabetically by author in a separate section of the references. The reviews permit an individual to read about another person's opinion of a book the authors have reviewed.

Selected Topics by Author

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Author Name		ь	C	đ	., e	f	g	h	i	1.	k	
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Chatterjee	X	X	diation.				X	`X `	· · · ·	X	X	
Cohen (1)	X	, X	X	X	X	X	X	X	X	X	X	3
Cohen (2)	X	X	X	X	́х́	X	X	X	X	X	x	2
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e = zero-order correlation

f = partial/semi-partial correlation

g = multicollinearity

h = dummy, effect and/or contrast coding

i = matrix algebra

j = residual analysis/outliers

k = variable selection methods

1 = covariance

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•			Orie	ntation	Presen	tation	Backgr	ound
Author	Year	Publisher	a	b	C	d	е	f
Belsley	1980	John Wiley	X	x	X		1	
Chatterjee	1977	John Wiley	x			. x		, x
Cohen (1)	1975	John Wiley	x		X	x		X
Cohen (2)	1983	John Wiley	X		x .	x .		. X
Draper (1)	1966	John Wiley	x		X	X		×
Draper (2)	1981	John Wiley	x		x	x	x	
Duna	1974	John Wiley	x			x		x
Edwards (1)	1976		x		x			x
Edwards (2)	1979	W.H.Freeman	: X		x			x
Fraser	1979	McGraw-Hill		X		X	x	
Freund	1979	Marc-Dekker	x			x		x
Goldberger	1968	MacMillan		X		X	X	,
Graybill (1)	1961	McGraw-Hill	x	x	x	x	X	
Graybill (2)	1976	Duxbury		x	ta ser e d	x	X	
Gunst	1980	Marc-Dekker	x		X	X	*****	X
Haitovsky	1973	Hafner		x		X	X	
Huang	1970	John Wiley		· X		X	X	•
Kelly	1969	SIU press	x		X			x
Kerlinger	1973	Holt,R & W	x		x	x		x
Kleinbaum	1978	Duxbury	X		X	x		x
Koerts	1969	Rotterdam	x	X	x		x	
Lewis	1978	SIU press	×			x	· · · ·	x
McNeil	1975	SIU press	X		x			X
> Pedhazur	1982	Holt, R & W	x		x			x
Plackett	1960	Oxford press		x		x	x	
Rao	1965	John Wiley		x		X	x	
Searle	1971	John Wiley		x		x	x	
Smillie	1966	Ryerson	x		x			x
Sprent	1969	Methuen		x	x	X		x
Ward	197 3	Prentice	x		x	X		X
Williams	1974	MSS corp.	x		x			x
Wonnacott	1981	John Wiley	x		X			X
Younger	1979	Duxbury	X		x	x		X

KEY: a = applied

b = theoretical

c = textbook

d = reference

e = matrix algebra

f = basic statistics

Table 3 indicates the <u>Flesch Reading Ease Scale</u> used to interpret the book score listed in <u>Table 4</u>. All books ranged from fairly difficult to very difficult which would be expected given the topics discussed. The readability measure does not take into consideration the numerous formulae, graphs, notation and mathematics. It does however provide some indication of readability for comparison among the books as well as a general indication of reading complexity compared to other types of reading material.

CONCLUS ION

The information provided permits comparisons among several books of multiple regression published over past years. Certain topics were indicated as appearing in the majority of the books. Most of the books reviewed emphasized an applied orientation with a basic statistics background requirement. Additional inquiry about certain books is possible by referring to the published reviews. The Flesch Ease Formula was used to compute a score on each book. The books reflected a difficult reading level comparable to scientific and academic text.

Most books had an outstanding feature which became apparent during the review process. For example, Belsley covered analysis of outliers and sources of multicollinearity. Chatterjee covered multicollinearity, autocorrelation and ridge regression extremely well. Cohen (1) and Cohen (2) had the widest range of topics covered and included one of the few discussions of power. Draper (1), Draper (2), and Qunst present the analysis of residuals/outliers and variable selection techniques the best. Edwards (1) and Edwards (2) afford an excellent introduction to linear regression with the presentation of different designs for analysis with dummy, effect, and contrast coding. Graybill (1) and Graybill (2) offer a broad coverage of topics at an advanced level using a matrix algebra

Table	3
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Reading Fase Scale	Description Ty	pical Magazine	Grade Level			
0 to 30	Very Difficult	Scientific				
30 to 50	Difficult	Academic	:			
50 to 60	Fairly Difficult	Quality				
60 to 70	Standard					
70 to 80	Fairly Easy	Slick-fiction	7th			
80 to 90	Easy	Pulp-fiction	6th			
90 to 100	Very Easy	Comics	5th			
100			4th			

Flesch Reading Ease Scale^a

^aAdapted from Flesch, 1948, p 230

Table 4

R.E. Score		Author	Description
16.5 18.1 23.1	-7	Cohen (2) Pedhazur Belsley	Very Difficult
32.6 32.9 33.1 33.8 35.1 35.9 35.9 36.8 38.2 39.5 42.7 45.4 46.2 46.5 48.5 48.5		Kleinbaum Haitovsky Gunst Huang Cohen (1) Goldberger Smillie Draper (2) Fraser Kerlinger Rao Sprent Wonnacott Koerts Dunn Ward	Difficult
50.7 52.0 52.7 53.1 54.4 54.9 55.9 56.6 56.7 56.9 57.2 57.2 57.2 57.4 60.0	- -	Draper (1) Chatterjee Lewis Searle Graybill (2) Freund Edwards (1) McNeil Williams Kelly Younger Graybill (1) Plackett Edwards (2)	Fairly Difficult

Readability Comparisons

Mark Leftman the start start and approach. Kelly, Lewis, and McNeil propose model formulation to test given A Starting of the set research hypotheses. Kerlinger and Pedhazur cover dummy, effect, and contrast coding well. Pedhazur additionally included a computer program on LISREL. Kleinbaum provides a broad coverage of all topics with excellent multivariate examples. Williams provides excellent examples on coding repeated measure designs. And finally, Younger provides computer applications using SAS, BMDP, and SPSS. Overall, selection of a specific book for classroom use is in the "eyes of the beholder", but this information should permit an alternative to experimentation or chance

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Books Reviewed

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      Efron, B. (1967). American Statistical Association Journal. 62.
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      Association. 67. 954.
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      (1970). Choice. 7. 717.
 Ward
      Burstein, L. (1976). American Education Research Journal. 13.
      296-301.
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Williams: none found

Wonnacott: none found

Younger

Askin, R. G. (1980). Journal of Quality Technology. <u>12</u>. 55-56. (1980). Choice. 17. 111.

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Ellinger, P. J. (1981). Use of Readability Formulas in the Formative Evaluation of Patient Education Materials. Unpublished master's thesis, Southern Illinois University. Carbondale, Illinois.

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MICROCOMPUTER-RESIDENT PROGRAM FOR THE ANALYSIS OF

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ABSTRACT

A microcomputer-resident program for the analysis of structural equations (PASE) has been designed to provide the causal modeler with all of the usually desired estimates of coefficients in recursive causal models. The program is interactive and self-documenting, and requires only a series of option selections by the user. Output includes all of the usual regression coefficients, plus total causal effects decomposed into direct and indirect causal effects.

Structural equation causal models provide a powerful aid to assist in the substantive interpretation of social and educational processes. Unlike straightforward regression analyses, structural equation analyses permit the measurement, not only of direct causal effects, but also of indirect causal effects through other, causally intervening independent variables (Finney, 1972). For example, it is now well understood that the primary reason father's occupational status is so closely associated with son's occupational status is not that sons directly inherit their father's status, but rather that sons of fathers with high status attain educations of a level that allow entry into occupations of higher status.

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Wolfle (1980), among others, showed how application of the basic theorem (Duncan, 1966) or first law (Kenny, 1979) of path analysis could be used to aid in the interpretation of the causal effects of one variable in a model on another. While the application of the first law of path analysis provides a useful aid in interpretation, in many cases its computation is tedious in practice. Alwin and Hauser (1975), followed by Wolfle (1983), showed how a series of relatively simple regression equations could be used to estimate the direct and indirect causal effects in a hierarchical structural equation model.

The present paper describes a microcomputer program designed to provide the causal modeler with all of the usually desired estimates of coefficients in recursive causal models, and also yields all total, direct,

and indirect causal effects implied by the hierarchical causal ordering of variables in the equation.

The microcomputer program accomplishes this goal by the simple expedient of requesting the user to supply information about the causal order of variables in the model, and with this information calculates a series of reduced-form equations from which the total causal effects of independent variables are stored in the computer's memory. The differences between the reduced-form coefficients, or total effects, and the fully specified, or direct effect, coefficients are the total indirect causal effects. Algebraic proofs of this relationship have been provided by Griliches and Mason (1972) and Wolfle (1983).

The microcomputer program described in this paper, PASE: Program for Analysis of Structural Equations, was designed to provide the causal modeler with all of the usually desired estimates of coefficients in recursive causal models. The program is user friendly in that it is interactive and self-documenting, and requires only a series of option selections by the usar. The program requires the input of a zero-order correlation matrix from either an existing file or the keyboard. Output Includes all of the usual regression coefficients, plus total causal effects decomposed into direct and indirect causal effects.

System Regulrements

PASE was written for an Apple II or Apple II Plus microcomputer that utilizes Applesoft BASIC. System configuration must be a minimum of 48K RAM, and one disk drive operating on DOS 3.3. The program provides support for, but does not require, a printer for hard-copy output.

The Analysis of Structural Equations

The most important advance in social research methodology in the past 15 years has been the introduction (Duncan, 1966) to the social sciences of causal modeling techniques first worked out over 60 years ago (Wright, 1921, 1925). On the one hand, this development has been important to social theory, for the techniques of causal modeling provide an explicit link between theory and the equations used to test the hypothesized relationships. On the other hand, while the estimation methods for structural equations implied by causal models are not new, the techniques have proven to be invaluable aids in the interpretation of social data. One of the most important of these interpretative aids in causal modeling is the decomposition of zero-order associations among variables into various causal components (see Wolfle, 1980).

A zero-order association may develop for one or all of three reasons. The association may be spurious; that is, it can develop because two variables, say X and Y, are related because they are both caused by a prior variable, Z, or a set of Z's. To the extent that the relationship between X and Y is spurious, that portion is called a noncausal component of the zero-order association. The remaining portion of the association between X and Y is causal, and is called the total effect. Total effects may in turn be decomposed into <u>direct</u> causal effects and <u>indirect</u> causal effects. Direct effects in recursive causal models are nothing more than partial regression coefficients of a variable regressed on all causes of it. The indirect causal effect that can be traced through

causally intervening variables. Such coefficients, both direct and Indirect, can be expressed in either standardized or unstandardized (metric) form; the latter are often preferred, because standardized coefficients are relatively unstable from sample to sample or across populations (Duncan, 1975; Kim and Mueller, 1976).

Users of structural equation techniques need to keep in mind, however, that the interpretations of causal effects are model specific. If the causal model is plausible, the variables within it credibly ordered and accurately measured, then the interpretations of effects within it are plausible. If these conditions are not met, however, then the interpretations based on faulty models are themselves faulty.

Program input and Output

A new computer program written for the Apple microcomputer, called PASE (Wolfle, 1982), provides a potentially useful tool for estimating hierarchical, recursive causal models. Because such models depend upon least-squares estimation procedures, PASE provides all of the usual regression coefficients. In addition, PASE provides estimates of total causal effects, and decomposes these into direct and indirect components.

PASE permits the input of new correlation matrices along with means and standard deviations. All data matrices can be saved to disk for future analyses. The program thus permits either the input of new matrices or the reading of previously saved data. Data matrices can be reviewed, corrected, truncated, or expanded to the maximum-sized matrix (17 variables) analyzable with the 48K memory limits of the compiled version of PASE.

Once the data have been input, reviewed if desired, changed if necessary, and saved to disk as recommended, the program prompts the user for the number of equations in the causal model. The program next asks the user to specify the dependent variable, followed by a list of the independent variables. The program next requests the user to specify the causal order among the independent variables. With this information, the program proceeds with the calculation of all regression coefficients, both standardized and metric, and decomposes these into direct and indirect causal components. (if one desires, the noncausal component of an association may be calculated by the simple expedient of subtracting the total causal effect from the zero-order association.)

The output of PASE has been organized for easy review. The output menu gives the user the option of reviewing the regression results, the regression ANOVA table, the R-squares among the independent variables, and the decomposition of causal effects. If desired, all of these results may be directed to a printer.

The regression results include all metric slopes, beta weights, standard errors, and t-ratios for the independent variables. The value of the intercept and the R-square for the regression are also included.

The ANOVA table includes the usual regression, residual, and total sums of squares, along with their associated degrees of freedom and mean squares. From these the F-ratio is calculated, and presented along with the standard error of estimate and the regression R-square.

The R-squares among the independent variables may be viewed. A high value among these suggests the presence of multicollinearity, which

if present causes regression coefficients to be unstable in the face of slight changes in the zero-order correlation coefficients (see Gordon, 1968). In addition, standard errors are often inflated, and highly correlated independent variables often (and implausibly) have regression coefficients of opposite sign (see, for example, Muffo and Coccari, 1982).

The table of causal decomposition presents the total effect of each independent variable, along with its direct effect and total indirect causal effect. If there are no intervening variables between the causal independent variablo and the caused dependent variable, then the total effect is the direct effect.

An Illustration

To illustrate the use of PASE, refer to the causal model illustrated in Figure 1. The model is based on some analyses presented in Duncan, Featherman, and Duncan (1972), and the data taken from Duncan (1968). Of particular interest in this model is the relationship between ability and earnings; what is the expected relationship between intelligence and earnings, controlling for social background, educational training, and occupational prestige?

There are three endogenous variables in the model; therefore, there are three equations to be estimated. Focusing attention on the equation for earnings, X(1), one would specify upon request by the program that variable 1 is dependent. The user will then be asked to specify the variable numbers of the causes of X(1); therefore, the user will input variable numbers 2, 3, 4, 5, and 6, since all other variables in the model are hypothesized to cause earnings.

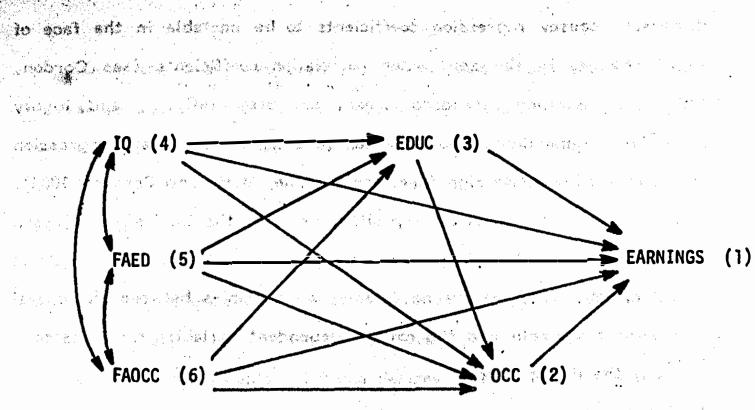


Figure 1. Path Diagram Representing Dependence of Earnings on Status Attainment, Intelligence, and Family Background.

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The program will next ask the user to identify the causal order of the independent variables. In this case, X(4), X(5), and X(6) occur simultaneously in a single causal block of exogenous variables. The user would therefore input variable numbers 4, 5, and 6 as constituting the variables in Block 1 (followed by the value 99 to terminate the Block listing). Educational attainment, X(3), is the single variable in Block 2, followed by occupational prestige, X(2), in Block 3. With this blocking information, the program proceeds with the calculation of the coefficients for X(1).

The regression results for this equation are shown in Table 1. These indicate by examination of the beta weights that the most important effect of earnings is the prestige level of the respondent's occupation, X(2). The relative effects of education and intelligence are less than half that of prestige, while the influence of father's education and occupation are statistically indistinguishable from zero.

The decompositions of causal effects for this equation are shown in Table 2. Examination of the total causal effects indicate that intelligence, educational attainment, and occupational prestige all have about equal total effects on earnings. The indirect effects indicate that about half of the total effect of education on earnings occurs indirectly through occupation; that is, those people with higher levels of educational attainment not only receive higher earnings <u>ceteris paribus</u>, but also tend to enter occupations of higher prestige which in turn lead to higher earnings.

Table 1. Regression Results

Dep	end	ent	Variable:	1

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•	Var.	B	Beta	St.Err	Т
- 3	2	.2625	.2625	.0381	6.8821
، ۳۰	3	.1069	.1069	.0423	2.5277
· •£.	4	.1013	.1013	.0344	2.9436
•	5	.0306	.0306	.0342	.8958
y area	6	.0183	.0183	.0348	.5263

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Variables:

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- 1 = 1964 earnings,
- 2 = 1964 occupation, 3 = education,

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- 4 = "early" intelligence, 5 = father's education,
- 6 = 'father's occupation.

FROM	TOTAL	DIRECT	INDIRECT
VAR. 4	.2273	.1013	.1261
VAR. 5	.0881	.0306	.0574
VAR. 6	.1032	.0183	.0849
VAR. 3	.2454	.1069	.1385
VAR. 2	.2625	.2625	0

Table 2. Decomposition of Causal Effects

(Standardized)

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Variables: See Table 1.

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The same may be said for the effects of intelligence on earnings. There is a direct causal effect of intelligence on earnings (the higher one's intelligence <u>ceteris paribus</u>, the higher one's earnings), but there is also a set of indirect effects wherein people of higher intelligence acquire higher levels of education and possess occupations of higher prestige, which also have positive effects on earnings. The combined direct and indirect effects of intelligence make it equally important to the explanation of earnings as is either education or occupational prestige.

In sum, PASE not only permits the causal modeler to examine the straightforward regression results, but, further, PASE also allows one to examine the decomposition of causal effects into their direct and indirect components. These latter examinations often prove to be very useful in revealing how causal effects are manifested in the model.

Availability of PASE

PASE is available from the author, College of Education, Virginia Polytechnic Institute and State University, Blacksburg, Virginia 24061. Please enclose one blank 5.25-inch, soft-sectored floppy disk compatible with the Apple disk operating system. A users' guide is also available; to cover duplication costs, please enclose a check in the amount of \$1.00 made out to VPI&SU College of Education.

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COMPARISON OF PROCEDURES FOR TESTING THE HYPOTHESIS OF A DIFFERENCE BETWEEN r 1 AND r 2 ISING INDEPENDENT AND DEPENDENT SAMPLES

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Introduction

Completion of correlation studies may require that the researcher test for significant differences between two independent correlations and/or between two dependent correlations. Solutions to the former problem may be found in many basic statistics books (Tate, 1965; McCall, 1970; Dayton, 1971; Minium, 1978). Procedures to test for a significant difference between dependent correlations have also been reported (Glass and Stanley, 1970; Hinkle, Wiersma and Jurs, 1979). Minium (1978) reported that there was no entirely satisfactory test of the difference between correlations from dependent samples, but it is not known whether he was familiar with the procedure presented by Hinkle, Wiersma and Jurs in 1979.

Method for Study One

suggested that differences between correla-Newman tions from both dependent and independent universes could be tested for significance using multiple linear regression (MLR). This application of the use of MLR had not been pre-viously demonstrated. While testing for a difference between r of independent universes appeared to be relatively uncomplicated using MLR, such was not the case when the test was applied to data from dependent universes. In the latter case repeated measurers were made, hence it was necessary to - 小型的物质系统 3 2 2 1 E 1 . S & include Person Vectors in the statistical models developed. Peddhazur, 1977 reported a procedure for inclusion of Person Vectors in MLR models, but no analogue procedure was given when the dependent variable was dichotomous. This paper presents such an analogue procedure and demonstrates its appropriateness.

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Results of using the procedure reported by Minium, 1978 to test for a significant difference between r and 1 r using independent samples and the procedures reported 2 by Glass and Stanley, 1970, and by Hinkle, Wiersma and Jurs, 1979 for testing the difference between r and r 1 2 using dependent samples were compared to results using the general MLR approach suggested below. Study of the outcome for the independent sample case was based upon a Monte Carlo approach in which 100 pairs of samples of 30 subjects each were taken from the Coleman Data Bank. The criterion vari-

Newman made the suggestion in planning the present paper.

able was sex (Y) and the predictor variables were GPA (X) and reading achievement (X). In the dependent $\frac{1}{2}$ case the same variables were used, but the subjects in sample 1 were the same subjects as those in sample 2. Using a Monte Carlo procedure, 100 samples of 60 subjects each were created from the Coleman Data Bank. When these subjects were considered to be in sample 1, a correlation (r) was calcu-1 lated between GPA and sex. When the same subjects were in sample 2 a correlation (r) was calculated between reading achievement and sex.

Comparison of Minium's Suggestion (z test) to MLR for Testing H_0 : $r_1 - r_2 = 0$, H_A : $r_1 - r_2 \neq 0$, A = .05 for Independent Sample Data.

Using a Monte Carlo procedure 100 pairs of independent samples were drawn. Correlations (r and r) were run 1between sex (Y) and GPA (X) and sex (Y) and reading 1 achievement (X). To determine if there was a significant 2 difference between r and r using the z test the 1 following formula was applied:

Formula One:

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$$z = \sqrt{\frac{\frac{z_{r_1} - z_{r_2}}{1}}{\sqrt{\frac{1}{n_1 - 3} + \frac{1}{n_2 - 3}}}}$$

Fisher's z equivalents were used rather than r values because the sampling distribution of the r values is likely to be skewed. Values of z obtained for the 100 pairs of

samples are reported in Table 1. Inspection of the z scores indicates that only two reached a magnitude greater than 1.96. Twice the null hypothesis was rejected with alpha set at the .05 level.

To determine if there was a significant difference between r and r with the same data using MLR, vari-1 2 ables X , X and Y were transformed into standard 1 2 scores to obtain common units of measurement. Using the following regression models, the hypothesis H : a = a0 1 2 (where a and a are partial regression weights) was 1 2 tested.

Full Model 1 $z_y = a_1 z_{11} + a_2 z_{12} + E_1$ VS $z_y = a_2 z_{12} + E_2$ VS $z_y = a_2 z_{12} + E_2$

(In standard score form z_{y1} represents sex, z_{x1} represents GPA, z_{x2} represents reading achievement and z_{x3} represents the predictor score regardless of whether the person came from sample 1 (s) or sample 2 (s); $z_{x3} = z_{x1} + \frac{1}{2}$ z_{x2} ; a represents the common slope for a and a .)

Full Model 1		с. Х.				
Model 1	z _y = a	$1^{2}x1 + 3$	$a_2^{z}x_2$ +	E ₁		
	^z y1 ₁		0	-	· .	· · · . · · ·
	^z y1 ₂	^z x1 ₂	0	-		ана ана м
⁹ 1	•	•	•	•		
	^z y1 ₃₀	• ^z x2 ₃₀	•	•		
	^z y1 ₁	0	^z x ² 1			: *
	^z y1 ₂	0	^z x2 ₂	-		
^s 2	•	•	•	•	-	•
	^z y1 ₃₀	0	^z x2 ₃₀	- *	•	a Sea
Restricted Model 2	Restr	iction:	9 7 0	•		
Model 2	z _y =	^a 3 ^z x3 [*]	^a 1 ^{= a} 2	2		
	^z y1 ₁	^z x1 ₁	· · ·	•		
s ₁	^z y1 ₂	^z x1 ₂		,		۰ •
·	• , •	•				•
:	^z y1 ₃₀	^z x1 ₃₀				
	^z y 1	^z x2 ₁				
	^z y1 ₂	^z x2 ₂		· · ·	· .	
⁹ 2	•	•		٩		
	^z y1 ₃₀	^z x ² 30				

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Results for Study One

Testing Model 1 against Model 2 will determine if a \neq a. The testing of Model 1 against Model 2 1 2 should give the same results as one would get by using formula one, the z test.

Reported in Table 1 are F values obtained by testing Model 1 against Model 2 for the 100 pairs of samples drawn (F critical for df₁'= 1, df₂ = 28, α = .05 = 3.34).

Only four of the F values computed when testing Model 1 against Model 2 exceeded the critical value or for this problem four times in a hundred a null hypothesis was rejected when alpha was set at .05.

When the z and F scores in Table 1 were compared, it was found that in 98 percent of the cases the same conclusion would have been drawn regarding the hypothesis H: r - r = 0. For two of the cases in which $0 \quad 1 \quad 2$ the F scores exceeded the critical value, this was also true of the z scores. Examination of cases 44 and 80 show the F scores exceeded the critical magnitude while the z scores narrowly failed to reach significance. (Critical z = 1.96, observed z scores were 1.88 and 1.86 respectively.)

Comparison Data for Independent Samples Testing the Hypothesis that $r_1-r_2=0$ Using MLR Vs Z Test

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Sample	R12	R ₁₃	F	Z			
1	0382	1445	.2275	.3906			
2	.1097	0822	1.0266	.7051			
3	.1659	.2998	.0562	4921			
4	1066	0515	.0213	2024			
5	.2092	.1476	.0341	.2264			
6	2958	2087	•2870	3202			
7	2195	2314	.0277	.0439			
8	.1568	.1993	.0074	1563			
9	0877	3084	.0018	.8110			
10	.2246	4637	7.8551	2.5292			
11	0876	.0219	.2493	4023			
12	3548	0749	.8628	-1.0285			
13	• 0987	.1240	.0018	0932			
14	2053	.0243	.7295	8433			
15	1435	1563	.0075	.0473			
16	0480	.1629	•5986	7749			
17	.0000	1785	.6556	.6557			
18	2496	.0925	1.8128	-1.2570			
19	.1861	0689	.8114	.9370			
20	1435	1249	.0067	0682			

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Table 1 (Continued)

Sample	R ₁₂	R ₁₃	F	Z
21	2176	3741	.2903	•5749
22	.0697	~. 0965	.3642	.6106
23	.1512	0695	.7133	.8108
24	4703	2006	1.6368	9910
25	. 02 05	.1436	.2862	4523
26	1161	3024	.3117	.6844
27	.0474		.1613	.4369
28	0955	2009	•2045	.3874
29	.1499	•,0943	• 0608	.2045
30		• 0636	1.0676	9725
31	.0724	2793	1.9379	1.2923
32	2372		.1464	. 498 4
33	3727	1552	.9003	7990
34	0693	.1093	.4751	6564
35	1826	1275	.1153	2021
36	.25 86	3983	7.2476	2.4136
37	.0553	.1837	.7674	4716
38	1205	1058	.0764	0540
39	1014	1708	.0305	.2550
40	1141	0091	.0876	3857
41	0957	.0036	.2218	3650
42	.1683	1274	1.4081	1.0867

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mple	R ₁₂ is in the second	R ₁₃	n an	2	
13	.0874	.2461	.7123	5830	
4	2835	.2293	4.6929	-1.8841	
15	.0879	2377	1.8545	1.1964	
16	2411	0628	.6567	6553	
17	.0388	.1021	.0292	2326	
8	0666	1132	.0020	.1710	
19	1985	.0159	.6105	7879	
0	1084	.3008	2.1849	-1.5033	
1	1756	2217	.0206	.1692	
.2	2905	3968	.3242	.3905	
• 3 • • •	-,2900	3334	.0202	.1592	
4	0976	.0574	.3046	~. 5696	
, 5	3273	.0132	1.6870	-1.2510	
, 6	0844	3070	.9242	.8181	
7	1926	1973	.2412	.0175	
• 8	.0219	.2391	1.3448	7980	
, 9	2871	2601	.0793	0992	
0	4133	2789	.287 6	4938	
1	.2297	.0214	.5132	.7654	
2	2553	0265	.4226	8410	
3	.0067	0732	.0108	.2935	

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Table 1 (Continued)

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Sample	**************************************	R13	F	
0822.**	in the the State of the			
64	1030	1145	.0104	.0424
65	1915	0339	.6764	5789
66	2034	1349	.0401	2517
67	1258	2431	.0679	.4310
68	3065	1268	.3624	6604
69	0125	1422	.2748	.4766
70	2968	.1366	2.6689	-1.5924
71	1137	 1915 ·	.0614	.2859
72	2050	1759	.0064	1069
73	0795	•0732	.2265	5611
74	.2440	2026	2.7359	1.6410
75	.1383	0037	.3197	.5215
76	.0074	0249	.0156	.1189
77	.0569	.0548	.0231	.0079
78	2532	3867	.4696	.4907
79	1658	.1182	.8270	-1.0434
80	3158	.1903	3.7560	-1.8594
81	2261	.0801	1.5966	-1.1251
82	- .2011	0749	.2890	4638
83	0156	.0234	.0150	1433
84	1663	0888	.1434	2845

Table 1 (Continued)

m ple	R ₁₂	R ₁₃	F	2 2
35 .	1278	.1714	.9152	-1.0993
6	1424	0176	.0748	4583
· 7	3278	.1275	3.1805	-1.6732
8	0852	.0105	.1551	3515
9	1290	0545	.1322	2736
0	0962	0521	.0419	1620
1	.0666	0978	.3745	.6041
2	1241	3654	.6108	.8869
3	.0000	1096	.7583	.4025
4	2129	.0429	.7594	9398
5	3984	.0154	1.8932	-1.5204
6	.0182	.1301	.0856	4111
1	.0736	.0088	.0604	.2383
9	2708	0433	1.1240	8360
•]	.0142	1267	. 3515	.5177
0	0968	1105	.0084	.0506

O: The F and Z values for the 100 samples were in agreement 98% of the time. Data for sample 44 and 80 showed the F with 1 and 28 df to be significant while the Z value narrowly fail to be significant. Critical value of F was 3.34.

<u>Comparison of the Glass and Stanley Procedure (z test) to the</u> <u>Hinkle, Wiersma and Jurs Procedure (t test) to the Newman Proce</u> dure (MLR) in Testing the Hypothesis $H_0: r_1 - r_2 = 0$,

 $H_A r_1 - r_2 \neq 0$ A = .05 for Dependent Sample Data.

Method for Study Two

Solution to the problem of testing for a significant difference between r and r when dependent samples 1 2 are used must take into account the lack of orthogonality by including the degree of co-variance between the two samples in the error term of the test. Results of solving this problem using the three procedures referred to will be reported below.

A Monte Carlo procedure was used to draw 100 pairs of dependent samples. Correlations were run between similar predictor (X) and criterion (Y) variables in each sample (r and r). The criterion variable (Y) was the 1 2 dichotomous variable sex. Predictor variables were GPA (X) and reading achievement (X). All data were 1 2 obtained from the Coleman Data Bank.

Formula 2, presented below, is the solution suggested by Glass and Stanley, 1970.

Formula 2 z =

$$\sqrt{(1-r^2_{xy})^2 + (1-r^2_{xz})^2 - 2r^3_{yz} - (2r_{yz} - r_{xy}r_{xz})^2} + (1-r^2_{xz})^2 - 2r^3_{yz} - (2r_{yz} - r_{xy}r_{xz})^2}$$

$$(1-r_{xy} - r_{xz}^2 - r_{yz}^2)$$

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Inspection of z scores obtained using Formula 2 and reported in Table 2 indicate that two of the 100 tests reached the critical value of 1.96. Thus, in only two cases was the null hypothesis rejected.

Formula 3 presented below is the solution suggested by Hinkle, Wiersma and Jurs, 1979.

Formula 3

$$t = \frac{(r_{xy} - r_{x2})}{\left| 2(1 - r_{xy}^2 - r_{xz}^2 - r_{yz}^2 + 2r_{xy}r_{xz}r_{yz}) \right|}$$

Inspection of t scores obtained by using Formula 3 and reported in Table 2 indicates that two of the 100 tests reached the critical value of 2.00 with df = 57. Thus, in only two cases was the null hypothesis rejected.

The MLR procedure used to test for a significant difference between r and r obtained from dependent samples involved the transformation of predictor and criterion variables into standard scores in order to obtain common units of measurements. The hypothesis that H : a = a = a (a represents the common slope) was $0 \ 1 \ 2 \ 3 \ 3$ tested by comparing the amount of variance accounted for by the following regression models. Values a , a , a are $1 \ 2 \ 3$ partial regression weights. Full Model 3 VS Restricted Model 4 Theoretical Models 3 and 4 $z_y = a_1 z_{x1} + a_2 z_{x2} + a_4 p_1 + \dots + a_{60} p_{63} + E_3$ VS $z_y = a_3 z_{x3} + a_4 p_1 + \dots + a_{63} p_{60} + E_4$

Peddhazur's conceptual approach for Models 3 and 4 where small ps are collapsed and designated as a large P. (See Peddhazur, 1977; Williams, 1977.)

 $z_y = a_1 z_{x1} + a_2 z_{x2} + a_4 P + E_3 VS z_y = a_3 z_{x3} + a_4 P + E_4$

(In standard score form z_{x1} represents GPA in sample 1 (s) z_{x2} represents reading achievement for the same persons in sample 2 (s), z_{x3} represents the predictor score regardless if the score came from sample 1 or 2; $z_{x3} = z_{x1} + z_{x2}$ and z_y represents the criterion variable sex; a is a partial regression weight; Ps represent person vectors used to account for the co-variance between the two dependent samples; a represents the common slope for the partial 3regression weights a and a .)

Below is a simulated numerical example to explain the procedure.

Full Model 3

Model	3	z ,	^a 1 ^z x1 +	^a 2 ^z x2 + a	$P + E_3$
Sub.	1	1	1	0	2.5
• · · ·	2	1	.5	0	1.2
s 1	3	0	3	0	5
	4	0	.7	0	1.6

	·'						าร์ มีนูรรับส
•	Sub.	1	0	0	1.5	2.5	
	•	2	0	0	.7	1.2	
	⁹ 2	3	1	0	2	5	
•		4	1	0	.9	1.6	· · · ·
Restr	icted	Model	4				· · .
Restr	ictio	n a ₁	= a ₂ =	^a 3	•	· .	
	Mode	14	^z y ^{= a} 3	$z_{x3} + a_4$	$P + E_{4}$		
•	Sub.	1	1	1	2.5		
•		2	1	.5	1.2	•	
s ₁	. 3	3	Ô	3	֥5		· ·
•	• -	4	0	•7	1.6		
• •	Sub.	1	0	1.5	2.5		
		2 •	0 ·	.7	1.2		
	⁸ 2	3	1	2	5	, , ,	
		··4 · · ·	na t ana ang	.9	1.6	, , , , , , , , , , , , , , , , , , , ,	

and the second part of the second second

Attention is directed to the procedure used to develop the person vectors. Model 3 represents the prediction of sex (z_y) by the standardized GPA (z_{x1}) , the standardized reading achievement score (z_{x2}) and a composite person vector (P). In the simulated model there are four subjects, two males and two females, each of whom is measured twice; once on GPA and once on reading achievement. The person vector is then computed by adding the GPA score of subject 1 to the reading

achievement score of subject 1; which in the simulated case sums to 2.5. Similarly for subject 2 one adds GPA to reading achievement and places the total 1.2 in the two positions of the person vector representing subject 2. This procedure is repeated until all persons are represented by a person vector, thus accounting for the covariance between the dependent samples.

Results for Study Two

Reported in Table 2 are F values obtained by testing Model 3 against Model 4 for the 100 samples drawn (F critical for df₁ = 2, df₂ = 57, $\dot{\alpha}$ = .05 is 3.17). Only two of the F values computed exceeded the critical value. Thus, for only two cases was the null hypothesis rejected with alpha set at .05.

When the z, t and F scores reported in Table 2 were compared, it was found that for the same two cases (58 and 62) the z, t and F test results were significant. It is, therefore, apparent that there was 100 percent agreement among the three procedures used.

Conclusion

To the extent that the approaches suggested by Minium, 1978; Glass and Stanley, 1970; Hinkle, Wiersma and Jurs, 1979 are valid, the use of multiple linear regression has been demonstrated to be a viable procedure for testing for a significant difference between r and r with both $1 \qquad 2$ dependent and independent data. Results using MLR were in

Comparison Data for Dependent Samples Testing the Hypothesis that $r_1-r_2=0$ Using MLR (F) Vs the Z Test Vs the t Test

Comparison Data for Dependent Samples Testing the Hypothesis that r1-r2= 0 Using MLR (F) Vs the Z Test Vs the t Test								
Sample	R ₁₂	R ₁₃	R ₂₃	F	Z	t		
1	1081	1019	5947	.0141	0273	027		
2	0049	0730	4437	.0600	.3109	.3033		
. 3	0685	1147	4848	.0200	.2094	.2059		
4	2429	.1196	4731	2.0612	-1.7047	-1.6438		
5	0427	1308	3334	.1386	.4213	.4125		
6	2337	0451	5143	•5423	8629	8587		
7	1578	1723	4318	.0002	•0680	.0680		
8	1443	0763	4686	.1262	3113	3071		
9	- 1229	2118	4150	.1868	.4203	.420]		
10	0698	1742	5636	•0899	.4648	.4625		
- 11	•0388	2175	5485	• 9222	1.1573	1.1320		
12	2294	.1421	5152	2.1572	-1.7221	-1.6558		
13	0391	0373	-,5848	.0003	0080	0078		
14	2290	0820	0279	.1535	8143	7983		
15	0909	0583	5034	.0285	1466	1438		
16	0975	1126	3899	.0074	.0709	.0697		
17	1112	0085	4717	.2554	4666	4559		
18	0344	.0036	4687	.0137	1719	1675		
19	.0208	0351	6242	.0499	.2406	.2344		

(Continued)

Sample	R ₁₂	R ₁₃	R ₂₃	F	Z	t
20	1637	.0648	6433	•6383	9927	966
21	1286	0124	5309	.1273	5187	507
22	.1160	1413	3697	.9576	1.2280	1.1880
23	1486	2208	3999	.1420	.3449	.3464
24	0933	.0405	5993	.2844	5828	5674
25	1714	.0376	5277	•5597	9417	9184
26	.1634	1349	5247	1.2434	1.3573	1.3096
27	1369	.0733	2734	.6172	-1.0346	-1.0046
28	.0417	1020	4548	.2704	.6571	.639
29	•0000	0934	3753	.0726	.4381	.4273
30	.0378	.2072	3869	.4074	8049	7919
31	1102	.2080	4317	1.5068	-1.5030	-1.451
32	 1122	.0247	3253	.3005	6559	6389
33	1519	.0089	5131	.2809	7245	7 0,84
34	1304	.1154	5743	.8748	-1.0910	-1.0559
35	.1613	2518	5863	2.4361	1.8883	1.8093
36	2672	.1156	4620	2.2825	-1.8183	-1.754]
37	1817	.1420	6488	1.5265	-1.4215	-1.3691
38	2039	2760	3791	.3468	.3548	.3639
39	1299	1480	3844	.0000	.0858	.0849
40	.0268	2342	4445	1.0715	1.2244	1.1973

Sample	R ₁₂	R ₁₃	R ₂₃	F	Z	t
41	0862	.0675	5479	.3316	.6814	6624
42	1134	1388	3589	.0068	.1214	.1197
43	1972	0269	5492	.4731	.7643	.7557
44	.0722	1105	4592	.5056	.8365	.8126
45	.1039	0568	4406	.3424	.7391	.7186
46	1462	.0208	5723	.3459	7379	7210
47	1396	.0224	4709	.3753	7395	7212
48	1728	1842	6266	.0000	.0502	.0521
49	1147	- •0868	5359	•0033	1245	1230
50	•0794	.0287	5166	.0303	.2263	.2212
51	0849	0789	3848	•0007	0280	0274
52	1164	1037	4168	•0085	0591	0582
53	1752	0219	5276	.2925	6898	6789
54	2124	0444	4226	•4027	7889	7785
55	•0896	0502	3565	.2555	.6 618	•6439
56	.1172	2122	5267	1.5152	1.5087	1.4563
57	1864	.1915	6415	2.0257	-1.6806	-1.6098
58	4106	.0709	4124	4.1282	-2.4361	-2.3888
59	0425	0898	3059	.0451	.2278	.2226
60	.0822	1186	5032	.6037	.9077	.8809
61	2270	.0985	3751	1.5038	-1.5736	-1.5216
62	- . 1543	.3160	6016	3.6028	-2.1797	-2.0931

(Continued)

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			1997 - 1997 -		· · · · · · · · · · · · · · · · · · ·	·
Sample	R ₁₂	R ₁₃	R ₂₃	F	Z	e e
63	.0251	1084	4283	.2451	.6155	.5996
	2392	0805	3373	.3583	7738	7665
65	2249	2041	4498	.0027	0986	1010
66	1786	0302	2974	•4226	7250	7099
67	1096	.1093	3481	.6796	-1.0474	-1.0157
68	0522	•0362	4824	.1157	3990	3885
69	1247	.1739	3957	1.3316	-1.4216	-1.3729
70	.0057	-,1372	6823	.3335	.6094	.5983
71	1623	•1125	3261	1.0284	-1.3374	-1.2939
72.	1065	1328	4727	.0680	.1202	.1190
73	.0209	.0120	4677	.0018	.0405	.0395
74	2765	0479	3575	•5527	-1,1157	-1.1045
75	1378	0768	3822	.7065	1.0143	.9844
76	0664	1420	3096	.1457	.3661	.3591
77	1462	1286	3292	.0466	0850	0840
78	2009	0959	5044	.3175	4800	4800
79	0960	2412	2796	.1335	.7244	.7171
80	1449	.0812	4936	.7979	-1.0293	9985
81	1217	1347	4893	.0955	.0591	.0587
82	0315	2108	5069	•3846	.8179	.8087
83	1265	.0388	4502	.4362	7589	7388
84	1869	.1276	5317	1.4378	-1.4325	-1.3818

(Continued)

Sample	R ₁₂	R ₁₃	R ₂₃	n de la F iele	2	t t
85	1404	.0383	6209	.5040	7773	7583
86	1376	0578	4992	.1262	3607	3552
87	2128	.1423	4939	1.7676	-1.6518	-1.5890
88	1787	0716	5486	.0233	4796	4772
89	0400	0773	6466	.0179	.1600	.1570
90	0700	.0425	5035	.1760	5043	4909
91	0341	2049	5194	.3524	.7750	.7667
92	0233	2465	4014	1.0884	1.0644	1.0487
93	0624	0942	3143	.0317	.1524	.1491
94	.1593	0670	5301	.7893	1.0181	.9892
95	1582	1596	3849	.0001	.0067	.0067
96	•0196	3220	6075	2.6945	1.5569	1.5625
97	2074	.0764	4796	1.1570	-1.3129	-1.2740
98	0962	0658	5215	• 0722	1357	1334
99	1564	.0108	5323	.3164	7495	7332
100	2481	.0265	4969	.9063	-1.2694	-1,2449

Note: The F, Z and t values for the 100 samples were in agreement 100% of the time. Degrees of freedom for the F and t values were $F_1=2$, $F_2=57$ and $df_t=57$. The critical value of F was 3.17 for t it was 2.00.

98 percent agreement with the procedure suggested by Minium (1978) for dependent data. For the two cases (44 and 80) where the MLR results did not agree with the more traditional procedure, the observed values just missed reaching the critical level, 1.88 and 1.86 respectively. When data from dependent samples were evaluated, there was 100 percent agreement among the procedures suggested by Glass and Stanley, 1970 (z test); Hinkle, Wiersma and Jurs, 1979 (t-test); and Newman (MLR).

The similarity in the results tends to support the use of all procedures tested. The writers, however, found the traditional tests (z and t) to be more cumbersome when a computer program for testing general linear models was available. In addition to the pragmatic consideration, a pedagogical advantage seems to exist when using MLR. Teaching students how to use the general linear model permits them to conceptualize more clearly what they are doing. This would be especially true for more naive students for whom application of the traditional models may be based entirely upon what appears to be unrealted statistical procedures. For the more sophisticated individual, MLR facilitates expression of the research question of interest in terms of general linear models without having to worry about a specific procedure to use for that particular problem.

Further, it is the belief of the authors that the general linear model approach to testing hypotheses is more apt to increase the ability of the researchers to ask questions that are of most specific interest to them; reducing the likelihood of their making a Type VI Error, Newman, I.; Deitchman, R.; Burkholder, J.; Sanders, P.; and Ervin, L. (1976) and Roll, S.; Hoedt, K.; and Newman, I. (1979). A Type VI Error is the inconsistency between the research question of interest and the statistical model being applied.

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APPLICATION OF JUDGMENT ANALYSIS TO INTERRATER COMPARISONS

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33. **2**63.

A nultiple regression method is presented for comparing the bases of two raters' judgments. This technique, which has been referred to judgment analysis or policy as capturing, is described for judgments of two nurses. In the example presented, judgments of future infant performance were derived from the nurse's scoring of infants' behavior the Brazelton Neonatal Behavioral Assessment Scale. on Brazelton dimension scores served as predictors of future 主战:自己反威府 小谋 被推进上。 performance in a test of differences between the policies (criteria) of the two nurse-raters. Sample data illustrate the technique but do not constitute a direct test of the since the two nurse's ratings were actually on two data If the ratings had been on the different sets of infants. 医水浴性肉病 内热镜 重教 大学主义学 same babies or identical samples of babies, the technique would have revealed, first, that the two nurses based their judgments primarily on ono Brazelton dimension, interactive processes; and second, that one nurse consistently rated the babies' future performance at a higher level than did the This technique has potential application to other nurso. evaluation of rating criteria for training of observers or judges and in other problem solving areas such as conflict resolution.

Subjective predictions of progress and objective assessments of behavior are frequently required in Bany projects. Consistency and accuracy of these このゆうず ちつゆう しゅごとせきえもめ ピュータを優良素量的 programs and observations are important issues in evaluating the judgments of different individuals or policies in relation 被走着想数据虚伪的 网络加坡式工 co patterns of attributes (Fisher, 1983; Most & Starr, When assessing these judgments several questions 1983) . often arise, for example, which of the many observations contributed the most to the overall judgment? ΟΓ, aore importantly, if more than one observer is involved, to what . (* ÷ ; ; ; extent did the raters rely on the same criteria as the basis and the second second for their predictions. I what a we have lebox, former and

paper presents a general statistical method for This comparing the observations and determining the bases of the judgments of two individuals. The method is applied to observations on the current status and judgments of future 1. 17 A A A capabilities of newborn infants. The observations were made by two nurses in the process of conducting the Brazelton BNBAS, (Brazelton, Veonatal Behavioral Assessment Scale, 1973). Judgment of the infant's future performance was made ifter completion of the BNBAS assessment. To illustrate the aethod, the nurse's judgments are treated as if they were cating the same infants. The two raters' judgments are then compared in terms of the regression weights associated with UNBAS dimension scores (Als, Tronick, Lester & Brazelton, 1977) derived from the original BNBAS observations. These scores represent the following dimensions: 1. Interactive

capacity to respond to social stimuli through processes: 你了了?"你们的<u>你</u>会说。 orientation, cuddling and consoling; 2. <u>Motoric processes:</u> YMAR Say (Marager Consultant and Constant of Constant ability to maintain good tone, control motor behavior and 0000 . . integrate actions; 3. <u>Organizational processes</u>: ability to 1 1 modulate states of consciousness in interactions with the environment primarily by shutting out aversive stimuli; and 4. Physiological reaction to stress: stability in response to stress. Dimensions 1-3 are scored as follows: 1 for the state of the state of the **dood** 2 for average, and 3 for worrisome or deficient 1.172 Dimension 4 is coded either 1 (good) performance. or O 1 21 1 B 33

(bad) . Classif all standard pages of a so that says and

The general model used as a basis for analysis of degree of rater similarity is a multiple linear regression using least-squares estimation of the regression weights,

WW YWW IU & H2X1+ ···· + H3X2 + ···· + Wk Xk + E/ - 19 a. **(1)** a 20 a. 网络香露白根带 囊骨下 相思,于比到了此是一副同下,他们还没身后,并且保持了比喻,一群并不一乎。" 可以从 where w₁, w₂, and w₁... w_k are least squares weights that minimize the squared errors in E. U is a vector of "1"s and the sector of the sector o X_1 and X_2 ... X_k are the K predictor vectors. The dependent ÷. set of judgments or ratings of the variable, Y, is the predictor data. Tho situations characterized by the regression approach outlined here is a variation of 8 technique called "policy capturing," (Christal, 1968a, **b**: christal 6 hottonberg, 1969; Ward, 1979). The combination of the regression weights applied to each variable is taken as defining the rater's "policy" with regard to 7, the dependent variable.

The general hypotheses to be tested are: "Does the policy used by one rater differ from that used by another?" and "If the two policies differ; do they differ by a constant amount?"

The models for interrater comparison are presented first followed by their application to sample BNBAS data.

METHOD

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Model Development

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The following regression equations were designed to test the Judgments of the two raters, Nurse 1 and Nurse 2, on the four BNBAS dimension scores. Each nurse's equation would take the general form,

Y nurse = function of (Dimension 1, Dimension 2, Dimension 3, and Dimension 4) + E (2)

thore Y is a nurse's judgment of an infant's performance. A similar regression equation is established for Nurse 2.

<u>Moial</u> 1. Model 1, which incorporates both nurses quations into a single model, takes into account the cossibility that wurse 1 makes ratings of infants that yield in equation (weights a_0 , a_1 , a_2 , a_3 , a_4) that differs from the corresponding equation (weights b_0 , b_1 , b_2 , b_3 , b_4) of urse 2. The equation is:

 $Y = a_0P1 + a_1(P1*D1) + a_2(P1*D2) + a_3(P1*D3) + a_4(P1*D4) + b_0P2 + b_1(P2*D1) + b_2(P2*D2) + b_3(P2*D3) + b_4(P2*D4) + E1$

(3)

「お」「合け」 きゆいきえる。 where I is the vector of future infant performance ratings from both nurses, D1 to D4 are the four BNBAS dimension scores, P1 is "1" for Nurse 1 and 0 otherwise, P2 is "1" for Nurse 2 and 0 otherwise, and E1 is the error in Nodel 1. In other words, the nurses are assumed to have based their predictions on two completely different policies. The least squares solution for Equation 3 will yield two sets of weights that might be different. Dimension 1 for Nurse 1 (P1*D1) has one weight (a_1) assigned to it, dimension 1 for a server all an and all server and because the server Nurse 2 (P2*D1) may have another weight (b,) assigned to it, 法法教教史 主語 and so on. Furthermore P1 is assigned one weight (a,) and 2011年7月1日至6月1日至6月 P2 may have another weight (b_0) .

Model 2. To test the hypothesis that the two nurses! predictions differed by a constant, restrictions are imposed on Model 1 to obtain Model 2, Equation 4. To illustrate this point, we would act as if the hypothesis is: when two nursos are presented with 10 babies and asked to nake 10 prodictions independently on those babios, tho predictions will dlffcr by a constant The amount. restrictions implied by the hypotheses of constant differences aro:

 $a_1 = b_1 = c_1, a_2 = b_2 = c_2, a_3 = b_3 = c_3, and a_4 = b_4 = c_4$ Substituting these restrictions in Model 1 gives Model 2.

医囊 你能够做了这些问题的,你是你的你的,你是你的你的,你不知道你的?""你是你看到了我跟我……" = $a_0 P1 + b_0 P2 + c_1 D1 + c_2 D2 + c_3 D3 + c_4 D4 + E2.$ (c,) for the two nurses, but that the nurses! judgments will differ by the constant value a - b -

<u>Model 3 assumes that the policies used by</u> Nurses 1 and 2 are identical. The restriction on Model 2 implied by this hypothesis is $a_0 = b_0 = c_0$. Substituting this restriction in Model 2 gives Model 3, Market 2000 2

The second states of the secon $Y = c_0 U + c_1 D1 + c_2 D2 + c_3 D3 + c_4 D4 + E3,$ (5) where U = P1 + P2, the Unit Vector containing a "1" in every element. Observe that this model has given up all information that distinguishes the two nurses.

Testing the Hypotheses.

After Hodels 1, 2 and 3 (equations 3, 4, and 5) have been devaloped, the questions of policy differences can be answered by comparing the R2's (squared multiple correlations) from the equations. The guestion, "If the two policies differ, do they differ by a constant amount?" can be answarph by dotermining if Rg is significantly larger than N.2. This comparison, Equation 6, ĺs made by calculating

$$F = \frac{(R_1^2 - R_2^2) / (n_1 - n_2)}{(1 - R_1^2) / (N - n_1)}$$
(6)

which is distributed as \underline{F} with degrees of freedom (df₁) = (n,

- n_2) and $(df_2) = (N - n_1)$; n_1 (=10) is the number of coefficients in Model 1, n_2 (=6) is the number of coefficients in Model 2 and N (=45) is the total number of ratings by both nurses. If the P-test is not significant we accept the restricted Model 2, that is the hypothesis that the differences between the two nurse's policies are constant is not rejected. The difference will be $(a_0 - b_0)$. In this case Model 2 would be adopted.

The next step in the analysis depends on the result of the comparison between Model 1 and Model 2. If we reject the constant difference hypothesis we conclude that the policies differ, and, therefore, Model 1 is appropriate.

Caller States

If we accept the constant difference hypothesis Model 2 is assumed, and to test that the policies are identical we compare Model 2 with Model 3 as in equation (7),

$$F = \frac{(R_2^2 - R_3^2) / (n_2 - n_3)}{(1 - R_2^2) / (N - n_2)}$$
(7)

where R_2^2 is compared to R_3^2 . If R_2^2 is significantly larger than R^2 , the null hypothesis $(a_0 = b_0 = c_0)$ is rejected and it can be concluded that the nurses differ in their ratings and the difference is constant. If the difference in the two R^2 is not significant, it is concluded that there are no differences between the nurses! judgments when expressed in terms of the four BNBAS dimensions.

Model Application.

Subjects and Procedure. Subjects were 45 infants who were seen at term as part of a larger study of metabolic derangements, neurophysiological functioning and behavior. Informed consent was obtained from parents and physicians prior to testing. Brazelton assessments for 25 of these infants were conducted by one rater, Nurse 1, and the remaining 20 by a second rater, Nurse 2. The same assistant recorded the scores during the BNBAS tests done by both nurses. After each test was completed, the test information was combined to form the four dimension scores (Als et al., 1977). Subsequent to the determination of the four dimension scores the nurses made a Judged Future Performance, JPP, for each infant. This JFP was scored as 0, 1, or 2, to correspond with predictions of below average, average, or above average future performance. No other explicit criteria were suggested.

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<u>Regults</u> The scores for the four dimensions resulting from the tost of the two nurses are in Table 1.

Table 1. Judged Future Performance (JPP), and Brazelton Neonatal Assessment Scale Dimension Scores From Two Nurses

Nurse 1			Nurse 2							
Case JPP D1	D 2	D 3	D4	Case	JF?	D 1	D2	D3	D4	
1 2 1	2	1	1	26	2	· -· 1	2	1	· . 1	
2 1 2	1	2	1	27	1	2	1	2	1	
3 1 1 1	2	· 2	- 1	28	* 1	2	2	2	1	
4 1 2	2	2 3	1	29	1	3	2	2	1	
845 A. 1	2		1	30	2	1	2	2	1	
6 1 1 7 2 1	1	2	1	31	1	2 3	2	3	0	
	2	1	1	32			2	2	1	
8 1 2 9 1 2	2 2	2 3		33	1	3	2	2	1	
10 1 1 2	2			34	1	3	2	2		•
	1	2 2		35		2	2	2		
	2	_	0	36	2	1	• •	2		
12 1 2 13 0 3	i: 1	2		37		2	2	2	- I	
13 0 3 14 1 1	•		0		2 - i	2		_	. 1	
	3 ⊵_2' '	3		39		· 3	2	2	1	
16 1 3	2	2	1	4 U S	1	2	2	2	1	
17 1 3	1	2		+ <u>+</u> ∉ = 42	1	1	2	2	1	
18 1 2	1	2	i	43	2	1	1	2	1	
19 0 3	3	2	• 1	44	2	2	2	2	1	
20 2 1	2	1	1	45	2	2	2	2	i	
21 1 21	2	* 1	1 2			-	£		• • • #	
2? 1 3		3	1	·						
23 20 1 2 3	2 2	22	1				- 11 - 1			
24 2 1	2	1	1							L
25 1 1	··1	1	· · ·			š	× • •			

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The four dimension scores, nurse identification and ratings of future performance were then entered into the models previously described. The results were $R_1^2 = 0.931$; $R_2^2 =$ 0.926; and $R_3^2 = 0.912$. The R² values were entered into the <u>F</u>-test formulas with the appropriate degrees of freedom. First, Model 1 was compared with Model 2 using Equation 6.

Test 1:
$$F_{(4,35)} = \frac{(0.931-0.926) / (10-6)}{(1 - 0.931) / (45-10)} = .597$$
 (8)

Test 1 (Model 1 compared with Model 2) was not significant. In light of this result, Model 2 was assumed where $a_1 = b_1 = c_1$, $a_2 = b_2 = c_2$, $a_3 = b_3 = c_3$ and $a_4 = b_4 = c_4$. Since Test 1 indicated that nurses' judgments differed by a constant amount, Model 2 was compared to Model 3 in Test 2, equation (9), using equation (7) above.

The <u>F</u> of 7.26 was significant at p < .05; therefore, the null hypothesis, that $a_0 = b_0 = c_0$, was rejected. While the expected nurses' ratings of future performance differed by a constant amount, the constant difference was not zero. The estimate of the actual difference was $a_0 = b_0 = 1.93 = 2.24 = -.31$ (see Table 2).

Riefor adi ozar, teradi	P	F-Value*	Prob-
Predictor		(df=1,39)	
P1-Nurse One P2-Nurse Two	1.93	2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2	ta t
D1-Interactive Processes	32	16.75	.0002
D2-Motoric Processes D3-Organizational	02	- 04	.8471
Processes D4-Physiological Peaction	22	3.45	•0708
to Stress	• 26	1.63	• 2098

Table 2. Regression Results for Model 2 (Equation 4)

*F-Values result from the (1,39 degrees of freedom) that the corresponding coefficient is equal to zero test

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Since the differences between ratings were constant (Test 1), we can conclude that the relationships between the four BNBAS scores and the judgments of Nurse 1 did not differ from the relationships of Nurse 2. But Test 2 indicated that even though differences were constant there was a significant difference between the level of ratings of the two nurses. Nurse 2 tended to give higher ratings (.31) than Nurse 1.

Since the nurses adid not actually rate the same infants it cannot be determined whether these results reflect actual differences in the nurse's policies or differences in the two sets of infants. In this example the relationship of the four PNBAS scores for the judgments was the same for the two nurses; therefore, it was of interest to examine each of the Four coefficients c, c, c, c, . Inspection of the Model 2 regression equation in Table 2 reveals that the two based their judgments primarily on dimension 1 nursas (Interactive Processes). This conclusion is based on the small probability (p = .0002) associated with the hypothesis that babies who have the same scores on Dimension 2, 3, and 4, but different Dimonsion 1 ratings will have the The probability of .07 expected JPP ratings. នងផម associated with the test on Dimension 3 indicates that Organizational Processes also may contribute to the judgment 3 37 5 process.

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DISCUSSION BOCODINTE ONLY BOOK

te car conclude that the relation of the sources part of the sources Bvaluation of judgments of behavior based on TA EL observations is a situation that occurs frequently. It is important not only to know on what bases and 🗉 how 6 consistently the observer is making judgments, but also whether judgments of different observers or raters have Techniques which address these questions are similar bases. demonstrated in Test 1 and Test 2, multiple regression models which have been described as rolicy capturing. This approach describes the set of variables or observations that best characterize a judgment. A start odd of astronomic to the start of the start o

- 带着警察会议的,这位要求,这次,一个人的事件的奇怪是一个人的,他不会地 One possible application of judgment analysis or policy capturing would be training programs where the goals are to evaluate and increase degree of intra- and inter-rater the policy or combination of independent reliability. Iſ viriables (observations), does not account for a significant proportion of the variance in the dependent variable, it can be inferred that the judgment of the observer is, to a large degrae based on information other than that contained in the producermined set of observations. In other words, the person is utilizing information not summarized in tho bohaviors represented by the values of the independent variables in the equation. For example, if the observer is instructo1 to make an assessment of an infant's future performance based on the results of the BNPAS, and the BNBAS values to not support or predict the JPP, it may be that

and the stimuly and grand the OF SOME other knowledge of the child's home environment unknown factor was entering into this judgment. In this situation, it may be necessary to retrain the observer to eliminate other than specified information or it may be more desirable to reconsider the factors in the equation. If two raters (judges or observers) differ in their rating criteria, the criteria of the rater whose judgments best approximate actual future performance can be adopted as the standard for others. These same considerations could be pertinent to questions of conflict resolution, both in refining the dependent variable (Nost & Starr, 1983) and as a way of describing how decisions are arrived at in problemsolving or negotiation settings (Fisher, 1983).

CONCLUSION

This technique can be a valuable aid for detecting implicit weightings of unknown variables which result in unexplained variance in judgments, and for standardizing judgments, that is, insuring that they are based on the same criteria.

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Julgement Analysis for Characterizing

Campus Ministry Function

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Abstract

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The purpose of the study was to investigate the judgmental policies of campus ministry held by campus ministers at state-supported universities when the campus ministers were grouped according to the campus minister's ministry group, years of personal campus ministry experience, size of the student body, campus minister's position at the school, and the campus minister's age by decade of birth. The ultimate goal of the research was to provide both clergy and laity with a clearer understanding of the role of campus ministry at state-supported universities. The questionnaire used in the study was developed using the critical incident technique. Supervisors of campus ministry were asked to list the three most important ministry goals or role functions or campus ministry at state-supported universities. The responses were tabulated and a 17-item questionnaire was formed. In order to determine reliability, a pilot test of the questionnaire was conducted. The subjects (N = 276) who participated in the study be responding to the questionnaire were campus ministers in ten ministry group affiliations at state-supported universities during 1982. They rated 17 goals of campus ministry and gave a rating to a program of campus ministry that would have the 17 goals as principal objectives. The Judgment Analysis technique was used and the campus ministers were found to be clustered in six judgmental areas related to ministry group.

The 1969 Wesley Foundation study found that clarification of ministerial roles and the search for self-image were among the greatest concerns of campus ministers (Underwood, 1969). Campus ministry has been in existence long enough to have a very large professional staff and a physical presence on hundreds of campuses (Johnson, 1979). Although this specialized ministry has produced several generations of practitioners and many generations of clients, it is still unable to define its role (Hammond, 1979).

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Lanagan (1979) suggested that both the university and the church are involved in determining the role campus ministry plans on campus. The university sees campus ministry as an academic or student life force and asks what preparation the campus minister should have to serve and assist the college or university in achieving its goals. The religious organization with which the campus minister is affiliated sees the campus ministry as a component which fosters a religious atmosphere in the University.

The purpose of this research was to develop purpose statements that could be identified by campus ministers as being relevant to campus ministry and analyze the purpose statements according to the campus minister's ministry affiliation, size of student body, and campus minister's age. The sets of purpose statements can be utilized to provide both clergymen and laity with a clearer understanding of the role of campus ministry at state-supported universities and to provide educational organizations affiliated with campus ministry

with direction in planning continuing educational opportunities for campus ministers.

Procedures

The critical incident technique (Flanagan, 1954) was used to develop the instrument used in this study. One-hundred seventy-one supervisors of campus ministry were asked to state what they considered to be the three most important goals of a viable campus ministry. The responses of the supervisors were tabulated and the most frequent responses were used as items (goals) on the insturment (see Table 1).

 Goals of Campus Ministry

 Number
 Statement of Goal
 Short Title

 1.**
 To assist students in developing
 Biblically based life

 Biblically based life goals and in the integration of these into the vocation of their choice.
 goals.

- 2.** To provide opportunities for Fellowship.
- 3.** To provide worship opportunities on campus.
- 4.* To develop student leadership.

5.*** To lead students and faculty to become involved in the local church.

- 6.*** To nurture students who are considering the religious profession as a vocation.
- 7.* To expand the vision of students to invest their lives in meeting the needs of a hurting world.

Worship.

Student leadership.

Involved in the local church.

Religious vocation.

Invest in hurting world.

Number

- 8.* To organize groups for study and action upon special concerns and problems raised in the university.
- 9.** To assist persons in their search for religious identity.
- 10.** To provide opportunities for study in doctrine, religious beliefs, and church (denominational) policy.
- 11.* To provide students with
 opportunities for personal
 ministry.
- 12.* To nurture students and faculty in faith development.
- 13.* To create an environment (organizational structure) in which students can grow in their faith.
- 14.* To develop a visible community of faith on campus.
- 15.** To provide pastoral counseling.

16.* To help students and faculty relate their work in academia and in the larger world beyond the campus.

- 17.* To enable the faith community on campus to be able to share their faith with others on campus while respecting the beliefs, values, and lifestyles of those other people.
- 18. Assuming that all the foregoing are principal objectives for a campus ministry program, how valid would you judge the overall goal of that ministry to be?

Organize for study and action.

Religious identity.

Study of religious topics.

Personal minstry.

Faith development.

Environment for growth.

Visible community of faith.

Pastoral counseling.

Relate faith.

Sharing of faith.

Overall rating of goals.

*	Factor 1:	Developmental Role of Campus Ministry
* *	Factor 2:	Supportive Rola of Campus Ministry
***	Factor 3:	Denominational Identity Role of Campus Ministry

Each item on the instrument was scored from one to five. An item 1.1.1.1 and the part of the second second THE MANY MORE received a score of 5 if the dimension was scored as being very important to campus ministry; 4 if the dimension was scored as being of more than average importance; 3 if the dimension was scored as CONTRACTOR OF being of average importance; 2 if the dimension was scored as being 1.11 of less than average importance; and 1 if the dimension was scored as being of little or no importance. COPOS STATUS SOUCH

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Construct validity of the instrument was investigated using factor analysis. Three factors (constructs) were found to exist and are indicated in Table 1. They were Developmental Role of Campus Ministry, Supportive Role of Campus Ministry, and Denominational Identity Role of Campus Ministry.

The instruments were then mailed to 500 randomly selected campus ministers serving at state-supported universities. The participants were selected from 3,427 campus ministers whose names appeared on mailing lists obtained from the headquarters of National Campus Ministry groups. There were 276 usable responses and Table 2 shows the ten groupings by ministry affiliation.

The sample consisted of 226 males and 50 females and were distributed among four age categories (see Table 2). Almost 64% of the campus ministers were less than 43 years of age. The sample was further categorized by the size of the student body at the institution where the campus ministry was located (see Table 2). Over 65% of the campus ministries were located at campuses having more than 9,000 students.

Table 2

Profile of Campus Ministers

Ministry Group	Number	Percent
Jewish Student Union	19	6,9
Southern Baptist Campus Ministry	54	19.6
Campus Crusade for Christ	34	12.3
The Navigators	13	4.7
Catholic Campus Ministry	36	13.0
Lutheran Campus Ministry	21	7.6
Presbyterian Campus Ministry	13	4.7
Inited Methodist Campus Ministry	29	10.5
Interdenominational	34	12.3
Episcopal Campus Ministry	23	8.3
Age		
æss than 33	87	31.5
13 to 42	89	32.3
13 to 52	66	23.9
Greater than 52	34	12.3
Student Body Size		
Less than 2,500 students	35	12.7
2,500 to 8,999 students	61	22.1
9,000 or more students	180	65.2
TOTAL	276	100.0

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Judgment Analysis (JAN) was utilized to identify the patterns by 中国大学性学、教育学校、多生学校、 which campus ministers make decisions about goals. The patterns were identified through the formulation of an association between the items on the instrument and an overall item. The strength of this association is reflected in the value of the multiple correlation coefficient (R). In this case the overall item represented an evaluation of all the goals which were presented to the campus ministers (Table 1). The JAN procedure gave an \mathbb{R}^2 (multiple \mathbb{R} coefficient squared) for each individual grouping of campus ministers 如果是你们要能能完成了你们的 體 and an overall R² for the initial stage of the procedure. The initial stage consisted of all the groupings when each one is treated as an individual system. Two judgmental groups were then selected by the procedure and combined on the basis of the homogeneity of their prediction equations. This resulted in the least loss in predictive efficiency of the procedure. The loss in predictive efficiency was measured by the drop in \underline{R}^2 between the two stages. The grouping continued until all of the groupings were combined into a single cluster. Har Berther Barris

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A determination of the number of different judgmental groups that are present can be made on the basis of the drop in \underline{R}^2 at the different stages of the JAN procedure. Ward (1962) and Ward and Hook (1963) suggested that a drop greater than .05 between successive stages represented too great a loss in predictability.

Results

Mean responses of the 276 campus ministers are shown in Table 3. Goals which were rated as most important were number 7 (invest in

MEAN RESPONSE SCORES FOR GGAL STATEMENTS"

		Goal Statements																	
Category	1	2	3	4 :	5	6	7	8	. 9	10	11	12	13	14	15	16	17	16	
Ministry Group:																2 . *	ā.		
I. Jewish	2.7	4.3	4.0	4.3	2.5	2.6	4.0	3.5	4.6	4.0	2.8	3.7	4.1	4.4	4.3	4.0	3.2	4.	
2. Southern Baptist	4.8	4.3	3.7	4.3	4.3	3.9	4.4	3.0	4.0	3.6	4.5	4.5	4.5	4.3	3.9	4.4	4.5	4.	
3. Campus Crusade for Christ	4.8	4.1	2.5	4.8	4.2	4.2	4.8	1.8	3.5	2.4	5.0	4.7	4.8	4.4	2.4	4.6	4.9	4.	
I. Navigators	4.7	.3.8	1.6	4.2	3.8	3.2	4.B	1.8	4.1	2.7	4.9	4.5	4.8	3.6	2.7	4.4	4. 8	4.	
5. Catholic	4.4	3.8	4.4	3.9	3.3	3.4	4.5	3.4	4.2	3.9	3.9	4.4	4.4	4. 3	4.2	4.1	3.9	4.9	
5. Lutteran	4.0	3.9	4.2	3.8	3.3	3.7	4.3	3.3	4.2	3.6	3.8	4.3	4.2	3.7	4.2	4.4	4.0	4.4	
7. Presbyterian	4.0	3.7	4.2	3.8	3.1	3.7	4.4	3.9	4.3	3.5	3.6	4.3	4.2	4.2	4.5	4.5	4.1	4.	
B. United Methodist	4.0	3.6	3.6	3.8	3.3	3.7	4.2	3.2	4.2	2.9	3.7	4.1	4.2	3.9	3.7	4.3	4.0	4.4	
J. Interdenominational	4.0	3.6	3.4	3.7	3.6	3.6	4.6	3.9	4.2	3.2	3.9	4.5	4.2	3.8	4.1	4.4	3.8	4.4	
C. Spiscopal	3.6	4.0	4.4	3.8	3.3	3.0	4.6	3.7	4.4	3.8	3.9	4.6	4.3	4.4	4.3	4.6	4.4	4.4	
Minister's Age:																			
< 33 years	4.5	4.2	3.2	3.5	3.9	3.9	4.7	2.5	4.0	3.0	4.6	4.5	4.6	4.3	3.3	4.4	4.6	4.9	
33-42 years	4.1	3.9	3.8	4.0	3.5	3.4	4.4	3.3	4.2	3.5	4.0	4.4	4.4	4.2	4.1	4.3	4.0	4.3	
43-52 years	4.2	3.9	4.1	4.0	3.4	3.7	4.4	3.7	4.3	3.6	3.9	4.5	4.3	4.0	4.1	4.4	4.0	4.9	
> 52 years	3.5	3.6	3.4	3.6	3.4	3.2	4.3	3.5	4.0	3.3	3.2	3.9	3.9	3.8	4.2	4.2	3.9	4.3	
Student Body Size:																			
<2,500 students	3.9	3.8	3.9	4.1	3.6	3.3	4.4	3.1	3.9	3.4	4.0	4.0	4.3	4.0	3.8	4.3	4.1	4.4	
2,500-8,999 students	4.2	4.1	3.6	4.1	3.7	3.8	4.5	3.2	4.3	3.3	4.1	4.5	4.4	4.2	4.0	4.5	4.1	4.	
29,000 studiets	4.3	3.9	3.6	4.1	3.5	3.6	4.5	3.2	4.1	3.4	4.1	4.4	4.4	4.1	3.8	4.3	4.2	4.4	
	4.2	3.9	3.6	4.1	16	3.6	4.5	• •	4.1								4.2	анан (19 1993 - 19 19	

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*Rounded to nearest tenth

hurting world) and number 13 (environment for growth). Each of these received an overall mean rating of 4.5. The campus ministers rated goal number 8 (organize for study and action) as having the lowest priority with an overall mean of 3.1. The Campus Crusade for Christ campus ministers gave as high as or the highest ratings of all groups for 10 of the 17 goals. The Jewish campus ministers gave as low as or the lowest ratings of all groups for 11 of the 17 goals.

Twelve of the 18 goals were given as high or the highest ratings of importance by the youngest group of campus ministers. The oldest ministers held the highest rating for only one goal, number 15 (pastoral counseling). Indeed, the oldest campus ministers had as low or the lowest ratings for 13 of the 18 goals.

Fourteen of eighteen goals were rated as high or higher by these ministers from medium sized schools than by either the ministers from schools with small or large student bodies. The small school ministers rated only one goal higher than the other two groups. That goal was number 3, i.e., to provide worship opportunities on Campus.

In an effort to determine the goal orientations of the three classifications; i.e., ministry group, student body size, and age, the data were submitted to Judgment Analysis technique (JAN). Characteristics of the campus ministers who evaluated ministry goals were illuminated by JAN which incorporates the strength of association between the ratings of the 17 individual ministry goals and the overall goal rating.

Table 4 demonstrates the judgment analysis system of regrouping

					AN ITERA												
	Stage						งวิบ	dge									<u>R</u>
MINISTRY GROUP	I	1	2	. 3	4		5		6		7		8		9	10	. 80
GROUP	11	1	2	3	. 4	et.	5	 1-	(6,	7)			8		9	10	. 80
	111	1	2 .	3	4		(5,	8)	1000 1000 1000 1000		(6,	7)		9		10	.78
	IV	1	2	3	4		(5,	8)			(6,	7,	10)		9		.76
	v	1	(2,	5, 8)	3		.4		(6,	7,	10)			9		•	.74
	VI	1	(2,	5, 8)	(3,	9)	-1	4		(6	, 7,	10)					.67
	VII	(1, 3	, 9)	(2,	5, 8)		4		(6,	7,	10)						. 59
	VIII	(1, 3	, 9)	(2,	5, 6, 7	8,	10)		4	j.							.49
	IX	(1, 3	, 4, 9		(2, 5,	6.7.	8,	10)	• •								.35
	x	(1, 2	, 3, 4	, 5, 6,	7, 8, 9,	10)					-						.19
MINISTER	I	1	2	3	4												.47
AGE	11	(1, 3)	2	4											·	.41
-	111	(1, 3)	(2, 4)												•	. 32
	IV '	(1, 2	, 3, 4)													.20
STUDENT	I	1	2	3													.31
BODY SIZE	11	(1, 2)	3													.29
	111	(1, 2	, 3)														.23

Table 4

AN ITERATIONS BY CLASSIFICATI

classifications of ministers. This process determines the groupings The second second 2 \$ \$ A who have made similar patterns in evaluating the goals. Thus in the first part of Table 4 the goals are analyzed by ministry groups. 2 8 A Starting with ten groupings of ministers the JAN procedure shows that group 6 (Lutheran) and group 7 (Presbyterian) were the most alike in & the tail the way that the ratings of the 17 individual goals related to the States at the states of the overall goal. This combination of campus ministers produced a a man a state the negligibly small reduction in \underline{R}^2 from stage 1 to stage 2. The \underline{R}^2 ANTRALM RESERVED . indicated the association between the 17 goals and the overall goal And the second second pe for each iteration. That is, the $\frac{R^2}{R^2}$ of .80 indicated that 80% of A CARACTER STATE the variability in the evaluation of the overall goals was accounted What Miles \$5 Attacks for by the 17 individual goals. The iteration process continued to e marke this has been been all a state combine ministry groups until a .05 decline in the $\frac{R}{R}^2$ was noted. ale the by efferred At this time six different groupings of campus ministries out of and the second second second second · , \$ the original ten were revealed. Groups 1, 3, 4, and 9 are singletons A Bring & Charles & Manager of States and having distinct characteristics by themselves, whereas 2, 5, and 8 The proving word - 1 were marged and 6, 7, and 10 were marged owing to the homogeneity of 10 AN their rating policies. No. Can at

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Using age as a means of classifying campus ministers (the second part of Table 4) four distinct ways of perceiving the subsidy of the individual goals to the overall goal of campus ministry appeared. The third part of Table 4 shows the campus ministers to have two composite policies with respect to student body size. Those campus ministers from small and intermediate size student bodies tended to have the same viewpoint concerning the contribution of individual goals to the overall while those from the largest schools were significantly different.

Discussion

The Southern Baptist, Catholic, and Methodist groups seemed to perceive all of the items as moderately associated with the overall goal of campus ministry. The goal showing the greatest contribution was number 6 (religious vocation) followed by 7 (sharing of faith).

Another composite of ministry groups combined Lutheran, Presbyterian, and Episcopal who also showed moderation on goal statements. The Presbyterians perceived goals 11 (personal ministry), 14 (visible community of faith), and 16 (relate faith) as being the most worthy dimensions of a campus ministry endeavor. While the Lutherans were very high on goals 7 (invest in a hurting world), 9 (religious identity), and 14 (visible community of faith), the Episcopals were very high on 16 (relate faith).

The other four campus ministry groups, the Jews, Navigators, Campus Crusade for Christ, and the Interdenominationals, all had very different perceptions of what a campus ministry should be. The Jews showed negative perceptions of goals 2 (fellowship), 4 (student leadership), and 6 (religious vocation) followed by negative perceptions of 10-13 (study of religious topics, personal ministry, faith development, environment for growth). All other goals seemed to make no contribution to the overall according to the perception of the Jewish ministry group. According to the Navigators goal 6 (religious vocation) has the highest priority followed by 5 (involved in local church), 4 (student leadership), and 14 (visible community of faith) for inclusion in a campus ministry program, whereas, goal 15 (pastoral counseling) was definitely not desired as a facet of a

ministry program. The Campus Crusade for Christ group had high perceptions for goals 5 (involved in local church), 10 (study of religious topics), and 15 (pastoral counseling) as being foundations of a campus ministry program, whereas, the interdenominational group showed high interest in goals 15 (pastoral counseling), 16 (relate faith), and 17 (sharing of faith). The interdenominational group showed little interest in the other goals in defining their campus ministry except for number 2 (fellowship) which they perceived as not being a part of a program.

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When the campus ministers were grouped by age the older personnel showed the strongest feelings about the components of a ministry program. They perceived the "lynch pine" to be composed primarily of goals 7 (invest in a hurting world), 11 (personal ministry), 13 (environment for growth), 15 (pastoral counseling), 16 (relate faith), and 17 (sharing of faith). The two middle aged groups (33 to 42 and 43 to 52) showed rather modest priority on most of the goals. The youngest of the campus ministers, however, perceived goal number 5 (involved in local church) as highest priority in a program followed by 1 (Biblically based life goals) and 14 (visible community of faith).

In the grouping according to campus population, ministers employed at small and intermediate sized campuses tended to have similar perceptions concerning the constituents of a campus ministry program. They also seemed to have the strongest perceptions overall, particularly wherein they rated goals 2 (fellowship), 3 (worship), and 7 (invest in a hurting world) as not being a part of the campus

ministry goal. However, these ministers rated goals 1 (Biblically based life goals), 16 (relate faith), and 17 (sharing of faith) as being most contributory. Campus ministers from larger campuses tended to be very moderate across the board, that is, they viewed all goals as being moderately contributory to an overall campus ministry goal.

Conclusions

The study seems to have revealed a consensus of priorities concerning the components of a campus ministry mission. These components are revealed according to ministry group, age of the campus ministers, and size of the student body at the institution where the campus ministers are employed. Evidence indicates that Southern Baptists, Catholics, and Methodists dominate the campus ministry movement. They revealed a moderation concerning the components of the campus ministry mission and seemed to view the campus ministry as an extension of the affiliated institution of higher learning. Evidence further suggests that Lutherans, Presbyterians, and Episcopal campus ministers viewed the goals from the standpoint of a more orthodox form of protestantism. The literature seems to indicate that these divisions tend to have more rituals and liturgy in their activities. The Lutherans seemed to view the campus ministry as a church functioning as a community within the campus, whereas the Presbyterians tended to emphasize the importance of personal faith in campus ministry. The Episcopals on the other hand seemed to underscore the idea that the campus ministry mission should support an applied religious philosophy.

That is, religion should address questions dealing with the way one should live in contemporary times and how one should decide about situational ethics.

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The Jews seemed to perceive very little social context within their campus ministry commission. They viewed the charge very differently from all other groups. Information suggests a sort of introspection about their approach. They were interested in pastoral counseling, individual religious identity and local church involvement in their campus ministry mission.

Church involvement in the student's life appeared to be a cornerstone of the Campus Crusade's ministry. The Navigators seemed to emphasize a religious leadership orientation with a social context. Results also suggest the Navigators as being organizers of leadership development. The Interdenominational group stressed individual student growth and sharing faith with other individuals.

When the sample was reclassified according to campus population, those campus ministers from small and intermediate size campuses seemed more interested in individual aspects of religious manifestations. Moreover, they were somewhat negative on fellowship and group Worship. Ministers from the largest campuses seemed more attentive to social programming but were moderately involved in all 17 of the goals.

Although the lack of a clear understanding of the role of campus ministry may be a problem in the field, it can be assumed that the campus ministers participating in the present study had definite judgmental policies of campus ministry and were consistent in expressing them.

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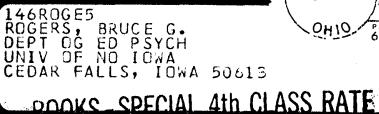
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